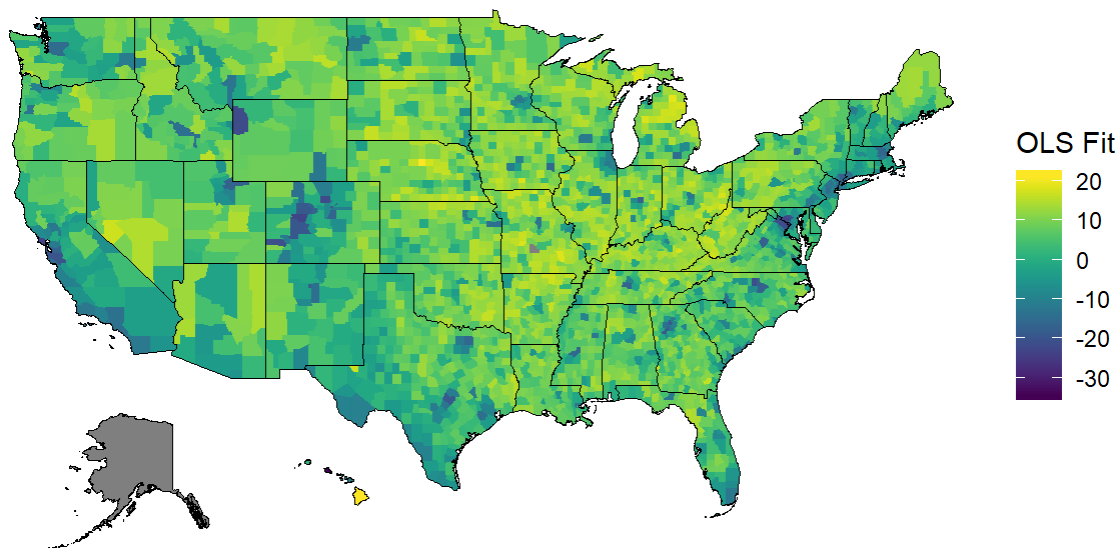


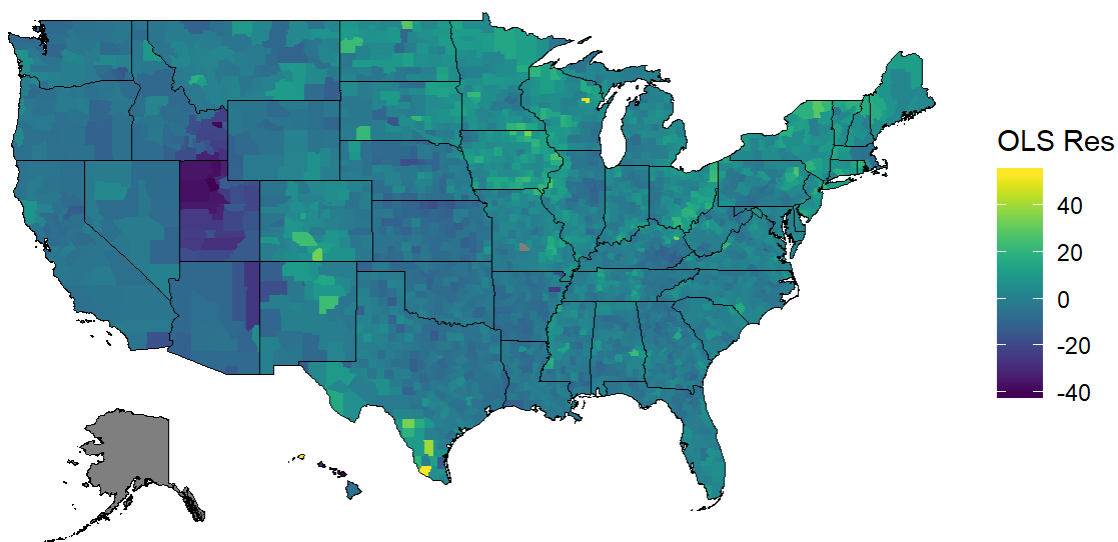
ST 433/533: Homework 5 Solutions

Non-spatial linear regression

Plot of fitted values



Plot of residuals



CAR model

Loading [MathJax]/jax/output/HTML-CSS/jax.js

Model description

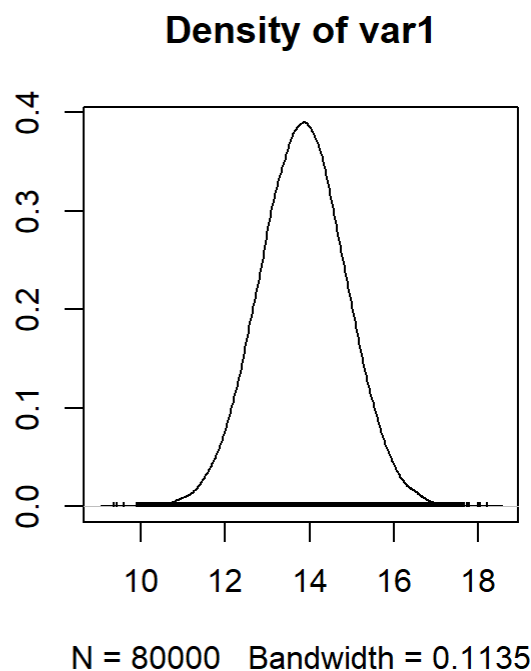
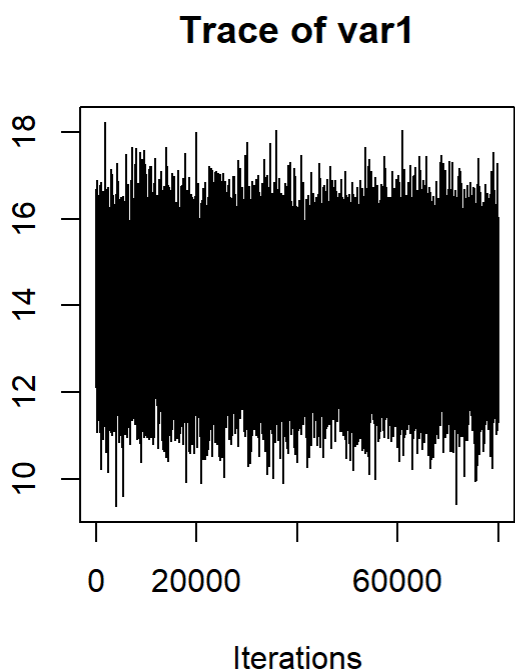
We are fitting the model $Y_i = X_i\beta + Z_i + \epsilon_i$, $i = \{1, 2, \dots, n\}$. The uncorrelated nugget error is $\epsilon_i \sim \text{Normal}(0, \tau^2)$. The spatial term is $Z = (Z_1, Z_2, \dots, Z_n)^T \sim \text{Normal}(0, \Sigma)$.

$Z_i | Z_{-i} \sim \text{Normal}(\rho Z_i, \sigma^2/m_i)$, where Z_{-i} is the collection of the $n-1$ other spatial terms and Z_i is the mean of Z_j over the m_i regions that neighbor region i . $\rho \in (0, 1)$ determines the strength of spatial correlation. σ^2 is the variance parameter. Using the Leroux parameterization, we have $\Sigma = \sigma^2[(1-\rho)I_n + \rho(M-W)]^{-1}$, where M is a diagonal matrix with m_1, m_2, \dots, m_n in the diagonal and W is the adjacency matrix.

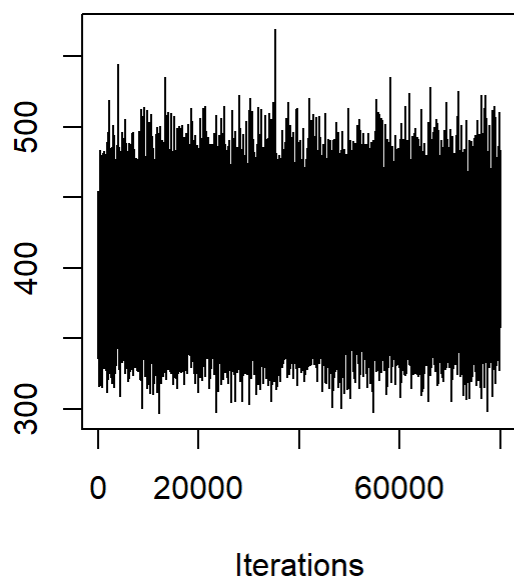
Y_i is the response and X_i is the set of covariates for county i . The estimates of β 's $\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_{10}$ are interpreted as the change in response Y_i with a unit change in the corresponding covariate, keeping the effect of other covariates fixed ($\hat{\beta}_0$ gives the estimate of Y_i when other covariates are unaccounted for).

MCMC convergence

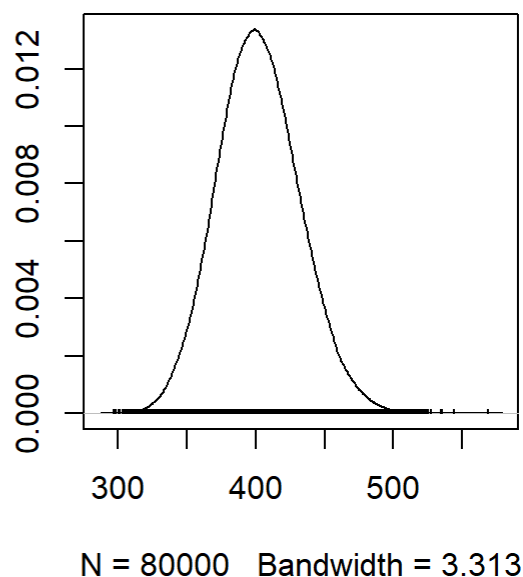
For σ^2 , τ^2 and ρ :



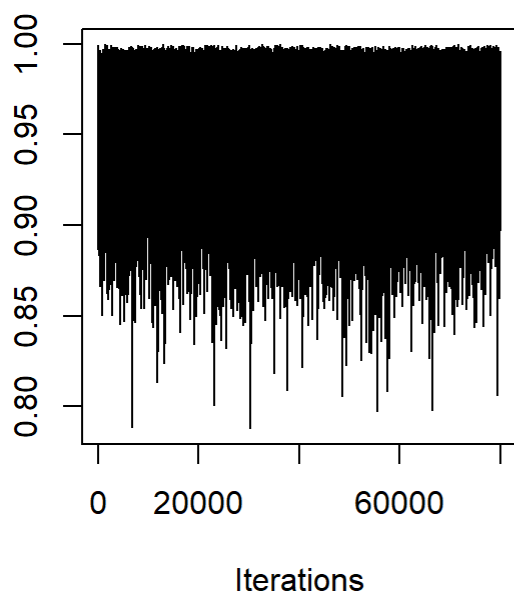
Trace of var1



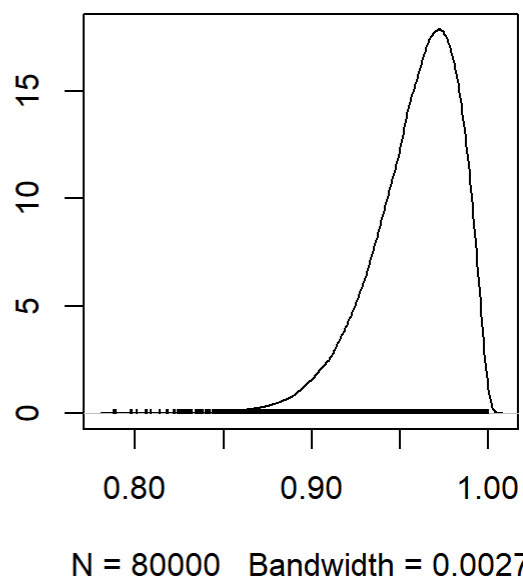
Density of var1



Trace of var1

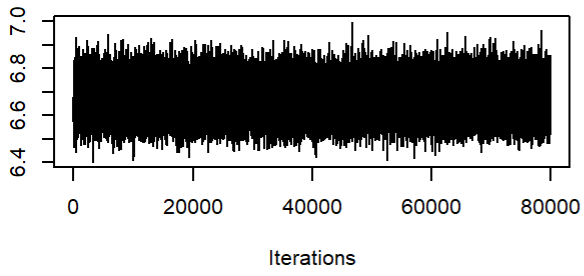


Density of var1

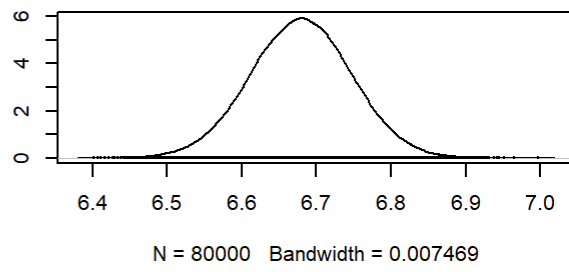


For $\beta_0, \beta_1, \dots, \beta_{10}$:

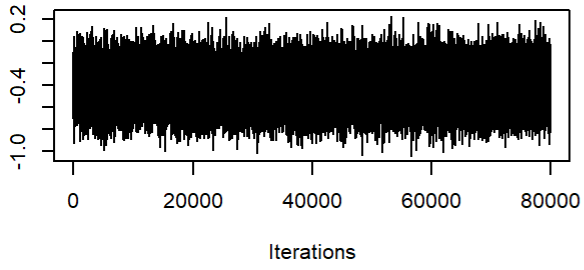
Trace of var1



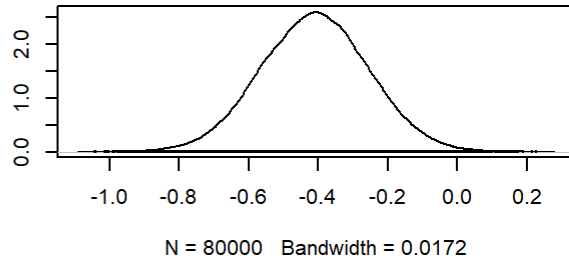
Density of var1



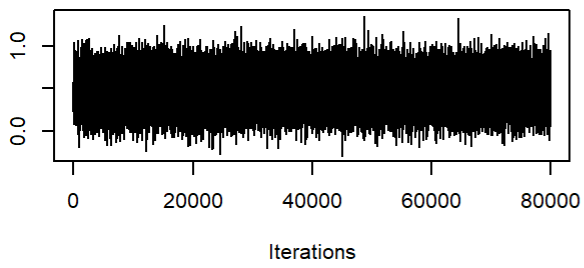
Trace of var2



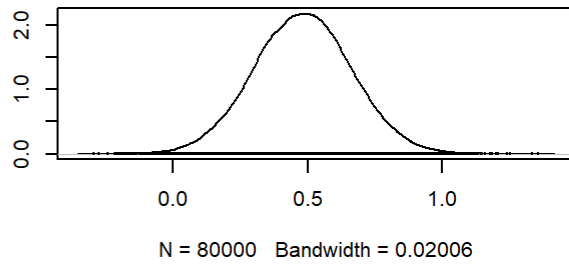
Density of var2



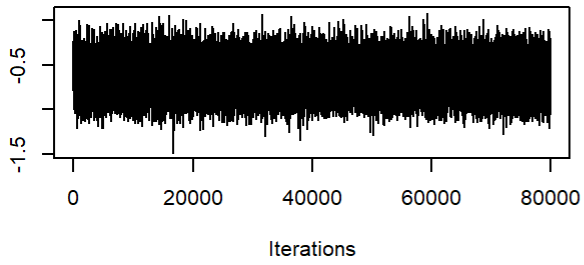
Trace of var3



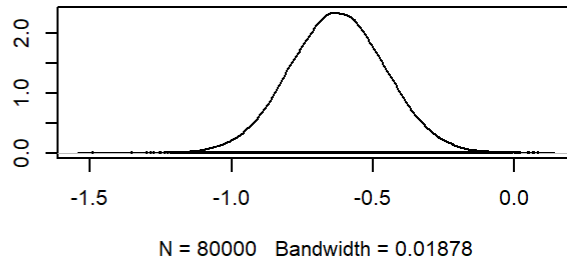
Density of var3



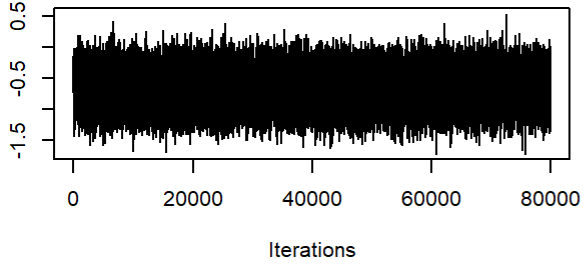
Trace of var4



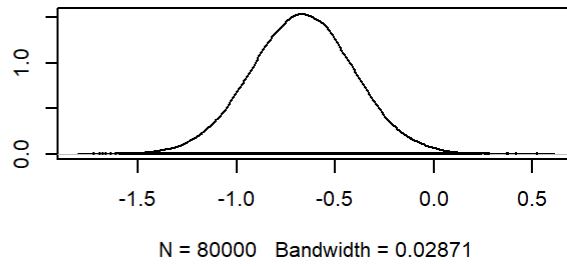
Density of var4



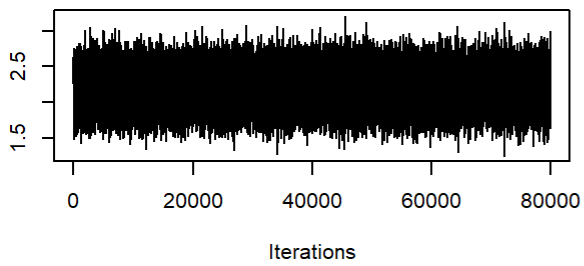
Trace of var5



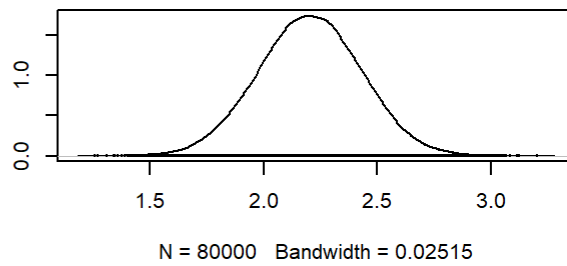
Density of var5



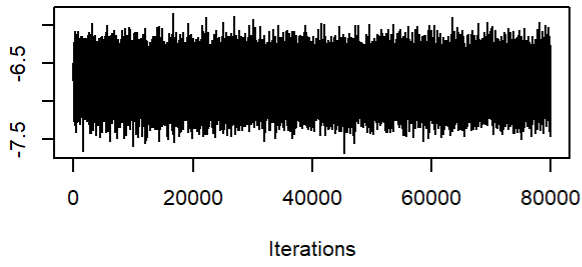
Trace of var6



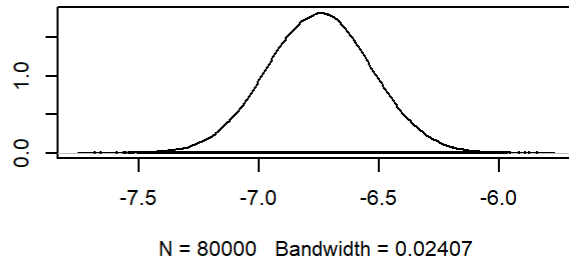
Density of var6



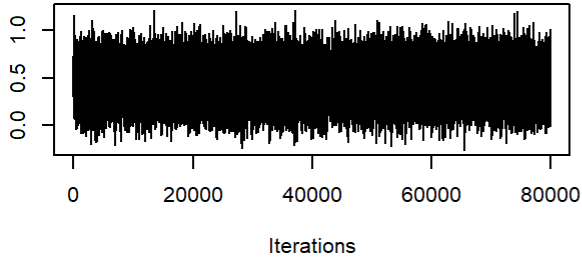
Trace of var7



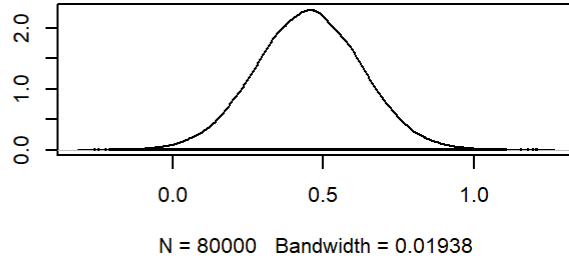
Density of var7



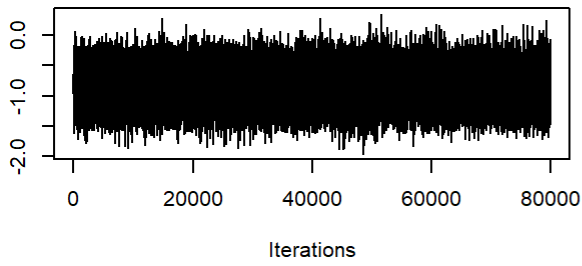
Trace of var8



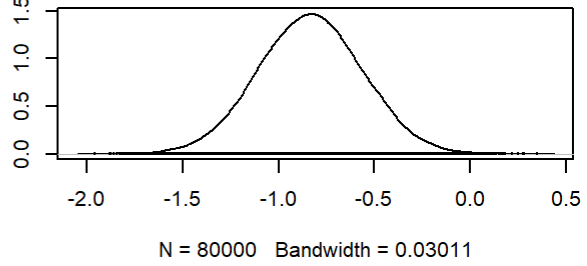
Density of var8

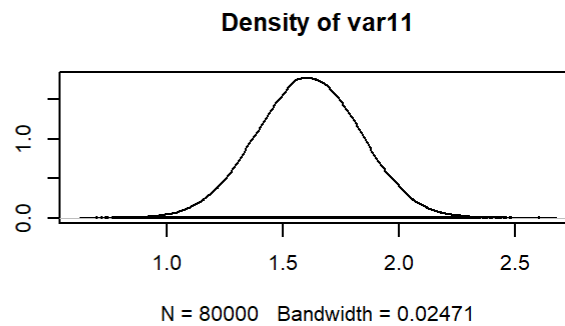
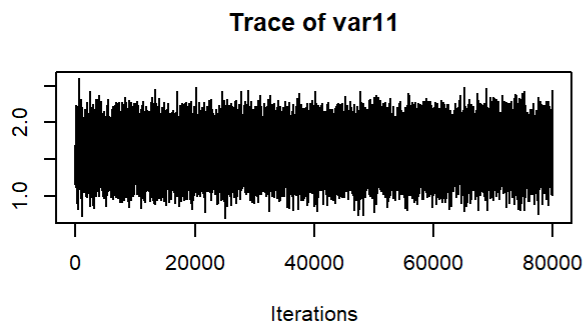
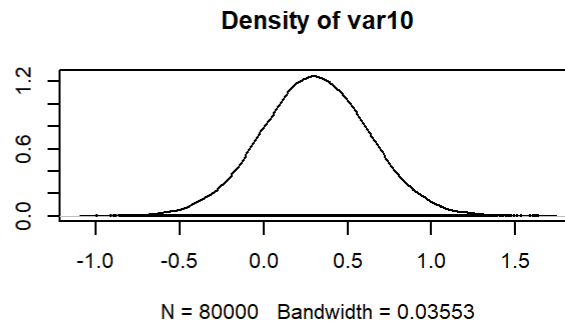
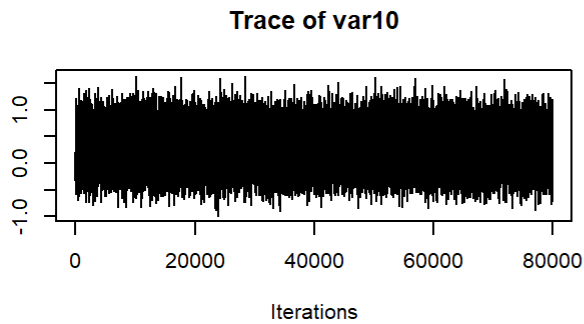


Trace of var9



Density of var9





For all the parameters, the MCMC has converged well.

Comparison

Non-spatial:

```
##
## Call:
## lm(formula = Y ~ X)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -40.268  -4.292  -0.703   3.870  52.981
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -1.088e+01  4.161e+00  -2.614  0.00898 **
## XPST120214  -2.693e-01  3.920e-02  -6.869  7.78e-12 ***
## XAGE775214   2.066e-01  4.508e-02   4.583  4.76e-06 ***
```

```

## XRHI225214 -1.095e-01 1.161e-02 -9.434 < 2e-16 ***
## XRHI725214 -1.526e-01 1.283e-02 -11.890 < 2e-16 ***
## XEDU635213 2.606e-01 3.697e-02 7.051 2.19e-12 ***
## XEDU685213 -7.169e-01 3.000e-02 -23.895 < 2e-16 ***
##XHSG445213 -1.572e-04 2.509e-02 -0.006 0.99500
##XHSG495213 -1.748e-05 3.007e-06 -5.815 6.70e-09 ***
##XINC110213 1.541e-04 3.166e-05 4.866 1.19e-06 ***
##XPVY020213 2.275e-01 4.416e-02 5.151 2.75e-07 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 7.482 on 3100 degrees of freedom
## (32 observations deleted due to missingness)
## Multiple R-squared:  0.4929, Adjusted R-squared:  0.4912
## F-statistic: 301.3 on 10 and 3100 DF, p-value: < 2.2e-16

```

Spatial:

```

##           Median      2.5%      97.5% n.sample % accept n.effective
## (Intercept)  6.6794  6.5461  6.8113  80000  100.0  80000.0
## XPST120214 -0.4098 -0.7141 -0.1062  80000  100.0  66283.7
## XAGE775214  0.4813  0.1263  0.8373  80000  100.0  65951.3
## XRHI225214 -0.6262 -0.9573 -0.2931  80000  100.0  52149.4
## XRHI725214 -0.6657 -1.1763 -0.1567  80000  100.0  28365.6
## XEDU635213  2.2024  1.7563  2.6444  80000  100.0  63395.3
## XEDU685213 -6.7484 -7.1715 -6.3239  80000  100.0  73890.3
##XHSG445213  0.4525  0.1076  0.7944  80000  100.0  72957.7
##XHSG495213 -0.8339 -1.3657 -0.3025  80000  100.0  49195.7
##XINC110213  0.3043 -0.3287  0.9364  80000  100.0  74666.8
##XPVY020213  1.6112  1.1732  2.0443  80000  100.0  78033.6
## nu2          13.8452 11.8403 15.8597  80000  100.0  28176.6
## tau2         401.1524 346.1194 463.4129  80000  100.0  29635.9
## rho          0.9628 0.8993  0.9935  80000  45.2  63094.4
##
##           Geweke.diag
## (Intercept)      -2.0
## XPST120214         0.2
## XAGE775214        -0.6
## XRHI225214         0.4
## XRHI725214         0.3
## XEDU635213        -0.5
## XEDU685213        -1.4
##XHSG445213        -1.6
##XHSG495213         1.9
##XINC110213        -0.8
##XPVY020213        -1.7
## nu2                1.5
## tau2               -1.4
## rho                1.2

```

The posterior means and sd for the coefficient estimates:

```
##           Post.Mean Post.SD
```

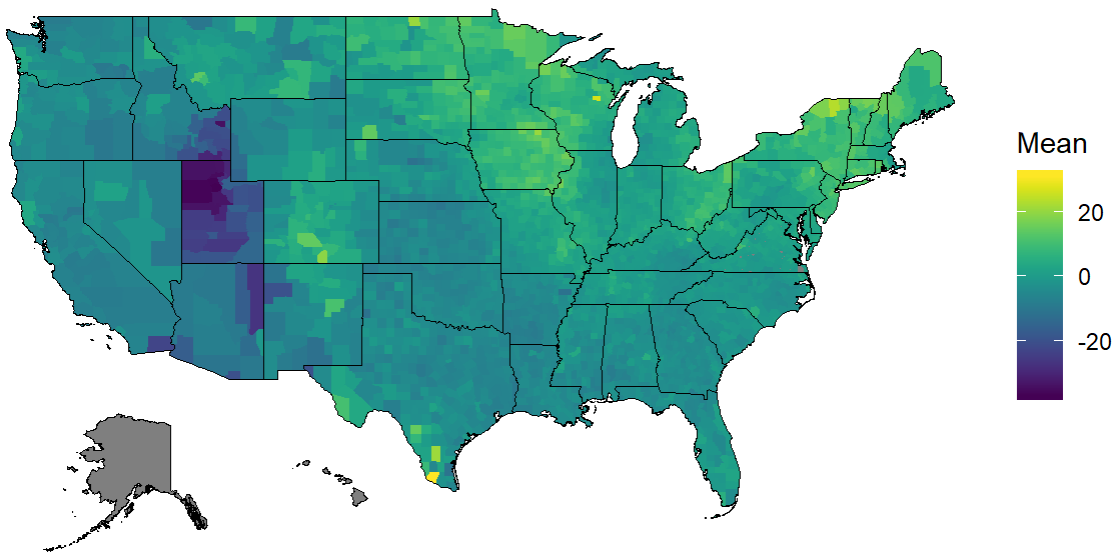


```
## var1      6.6791  0.0674
## var2     -0.4102  0.1552
## var3      0.4810  0.1810
## var4     -0.6256  0.1694
## var5     -0.6665  0.2590
## var6      2.2020  0.2269
## var7     -6.7486  0.2172
## var8      0.4520  0.1749
## var9     -0.8339  0.2716
## var10     0.3051  0.3208
## var11     1.6109  0.2230
```

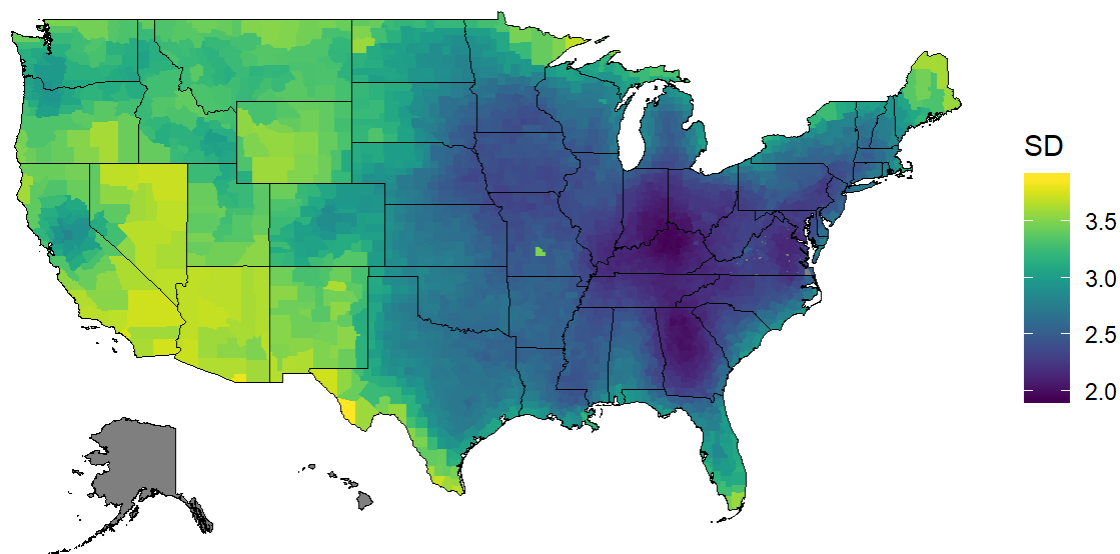
There is some difference when we compare the estimates from non-spatial and spatial model fits. The standard errors are somewhat larger for spatial model.

Spatial random effects

Posterior Means



Posterior SD



From the plots, the random spatial effect estimates are higher along the US-Canada border, a few parts of central US and Texas and a little part on the upper part north eastern border. Estimates have low values in and around Utah. The standard deviations are high on the south western part, in some parts of the lower boundary of Texas. Overall, there is lower sd in the Eastern part. A missing covariate could be: Employment rate.

Codes

```
library(tidyverse)
library(housingData)
library(fields)
library(readr)
library(parallel)
library(choroplethr)
library(viridis)
library(CARBayes)

load("election_2008_2016.RData")
county_adj2010 <- read_csv("county_adjacency2010.csv")

#Removing missing values
miss <- is.na(Y)
Y1 <- Y[!miss]
X1 <- X[!miss,]
fips1 <- fips[!miss]
all_dat1 <- all_dat[!miss,]

#Creating coordinates from fips
f <- as.numeric(paste(geoCounty[,1]))
```

```

s <- matrix(NA,nrow(all_dat),2)
for(i in 1:nrow(all_dat)){
  these <- which(fips[i]==f)
  if(length(these)==1){
    s[i,1] <- geoCounty[these,4]
    s[i,2] <- geoCounty[these,5]
  }
}

#Distance adjacency

d <- rdist.earth(s) # Distance in miles
d[is.na(d)] <- Inf # Set Alaska and HI to be infinitely far away
diag(d) <- Inf # Make sure counties don't neighbor themselves
nearest <- apply(d,1,min)
D <- max(nearest[nearest<Inf])

ADJ2 <- ifelse(d <= D,1,0)
#rowSums(ADJ2)

#Linear regression
lreg <- lm(Y ~ X)

#Residuals
res <- lreg$residuals

#Fit
fit <- lreg$fitted.values

#Define a function to make county maps using viridis colour
county_plot<- function(fips,Y,main="",units="",limits=NULL){
  library(choroplethr)
  library(ggplot2)
  library(viridis)

  temp <- as.data.frame(list(region=fips,value=Y))
  choro <- CountyChoropleth$new(temp)
  choro$title = main
  choro$set_num_colors(1)
  choro$ggplot_polygon = geom_polygon(aes(fill = value), color = NA)
  choro$ggplot_scale = scale_fill_gradientn(name = units, colours = viridis(32))
  suppressWarnings(choro$render())
}

county_plot(fips1,res,main="Plot of residuals",units="OLS Res")
county_plot(fips1,fit,main="Plot of fitted values",units="OLS Fit")

X <- scale(X) # Scale covariates
X[is.na(X)] <- 0 # Fill in missing values with the mean

```

```
W <- ADJ2

# Remove counties with no neighbors
junk <- rowSums(W)==0
Y <- Y[!junk]
X <- X[!junk,]
W <- W[!junk,]
W <- W[,!junk]
fips <- fips[!junk]

tick <- proc.time()[3]
modelCAR <- S.CARleroux(Y~X, family="gaussian", W=W, burnin=200000, n.sample=1000000, thin=10,
verbose=TRUE)
tock <- proc.time()[3]
(tock-tick)/60 # time in minutes

nu2 <- modelCAR$samples$nu2
tau2 <- modelCAR$samples$tau2
rho <- modelCAR$samples$rho
beta <- modelCAR$samples$beta
Zest <- colMeans(modelCAR$samples$phi)
Zsd <- apply(modelCAR$samples$phi, 2, sd)
Yfit <- modelCAR$samples$fitted.values

#MCMC convergence

plot(model1$samples$nu2)
plot(model1$samples$tau2)
plot(model1$samples$rho)
plot(model1$samples$beta)

#Random spatial effects
county_plot(fips, modelCAR$summ$Zest, main="Posterior Means", units="Mean")
county_plot(fips, modelCAR$summ$Zsd, main="Posterior SD", units="SD")
```